

# On trigonometric sums with random frequencies

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## Abstract

We prove that if  $I_k$  are disjoint blocks of positive integers and  $n_k$  are independent random variables with uniform distribution on  $I_k$ , then

$$N^{-1/2} \sum_{k=1}^N (\sin 2\pi n_k x - \mathbb{E}(\sin 2\pi n_k x))$$

has, with probability 1, a mixed Gaussian limit distribution relative to the interval  $(0, 1)$  equipped with Lebesgue measure. We also investigate the case when  $n_k$  have continuous uniform distribution on disjoint intervals  $I_k$  on the positive axis.

## 1 Introduction

Salem and Zygmund [7] proved that if  $(n_k)$  is a sequence of positive integers satisfying the Hadamard gap condition

$$n_{k+1}/n_k \geq q > 1 \quad (k = 1, 2, \dots) \quad (1.1)$$

then the sequence  $\sin 2\pi n_k x$ ,  $k \geq 1$  obeys the central limit theorem, i.e.

$$N^{-1/2} \sum_{k=1}^N \sin 2\pi n_k x \xrightarrow{d} N(0, 1/2) \quad (1.2)$$

with respect to the probability space  $(0, 1)$  equipped with Borel sets and Lebesgue measure. Here the exponential growth condition (1.1) can be weakened, but as Erdős

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[3] showed, there exists a sequence  $(n_k)$  growing faster than  $e^{\sqrt{k}}$  such that the CLT (1.2) fails. On the other hand, using random constructions one can find slowly growing sequences  $(n_k)$  satisfying (1.2). Salem and Zygmund [8] proved that if  $\xi_1, \xi_2, \dots$  are independent random variables on some probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  taking the values 0 and 1 with probability  $1/2 - 1/2$  and  $(n_k)$  denotes the set of indices  $j$  such that  $\xi_j = 1$ , then with  $\mathbb{P}$ -probability 1, the CLT (1.2) holds. For this sequence  $(n_k)$  we have  $n_k \sim 2k$  and by the theorem of "pure heads" we have  $n_{k+1} - n_k = O(\log k)$ . Berkes [1] showed that if  $\mathbb{N} = \cup_{k=1}^{\infty} I_k$  where  $I_1, I_2, \dots$  are disjoint intervals of positive integers such that  $|I_k| \rightarrow \infty$ , and  $n_1, n_2, \dots$  are independent random variables on some probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  such that  $n_k$  is uniformly distributed on  $I_k$ , then with  $\mathbb{P}$ -probability 1,  $\sin 2\pi n_k x$  satisfies the CLT (1.2). Thus, given any positive sequence  $\omega_k \rightarrow \infty$ , there exists an increasing sequence  $(n_k)$  of positive integers such that  $n_{k+1} - n_k = O(\omega_k)$  and  $\sin 2\pi n_k x$  satisfies (1.2). In [1] the question was raised if the CLT (1.2) can hold for any sequence  $(n_k)$  with  $n_{k+1} - n_k = O(1)$ . Bobkov and Götze [2] showed that the answer to this question is negative, and in particular, if in the construction in [1] we choose  $|I_k| = d$  for  $k = 1, 2, \dots$ , then with probability 1, the limit distribution of  $N^{-1/2} \sum_{k=1}^N \sin 2\pi n_k x$  is mixed normal. On the other hand, Fukuyama [4] showed, using another type of random construction, that for any  $0 < \sigma^2 < 1/2$  there exists a sequence  $(n_k)$  of integers with bounded gaps  $n_{k+1} - n_k$  such that (1.2) holds with a limiting normal distribution with variance  $\sigma^2$ . The purpose of the present paper is to return to the random models in [1], [2] and investigate the case of constant block sizes  $|I_k| = d$ , allowing arbitrary gaps between the blocks. We will prove the following result.

**Theorem 1.** *Let  $I_1, I_2, \dots$  be disjoint blocks of consecutive positive integers with size  $d$  and let  $n_1, n_2, \dots$  be a sequence of independent random variables on a probability space  $(\Omega, \mathcal{A}, \mathbb{P})$  such that  $n_k$  is uniformly distributed over  $I_k$ . Let  $\lambda_k(x) = \mathbb{E}(\sin 2\pi n_k x)$ . Then  $\mathbb{P}$ -almost surely*

$$\frac{1}{\sqrt{N}} \sum_{k=1}^N (\sin 2\pi n_k x - \lambda_k(x)) \xrightarrow{d} N(0, g) \quad (1.3)$$

over the probability space  $((0, 1), \mathcal{B}, \lambda)$ , where  $\mathcal{B}$  is the Borel  $\sigma$ -algebra in  $(0, 1)$ ,  $\lambda$  is the Lebesgue measure,

$$g(x) = \frac{1}{2} \left( 1 - \frac{\sin^2 d\pi x}{d^2 \sin^2 \pi x} \right) \quad (1.4)$$

and  $N(0, g)$  denotes the distribution with characteristic function  $\int_0^1 e^{-g(x)t^2/2} dx$ .

Here  $g \geq 0$  and  $N(0, g)$  is the distribution of  $\sqrt{g}\zeta$ , where  $\zeta$  is a standard normal random variable on  $(0, 1)$ , independent of  $g$ . Clearly,  $N(0, g)$  is a variance mixture of zero mean Gaussian distributions.

Note that  $\sum_{k=1}^N \lambda_k(x) = \mathbb{E}(\sum_{k=1}^N \sin 2\pi n_k x)$  is the averaged version of  $\sum_{k=1}^N \sin 2\pi n_k x$ , a nonrandom trigonometric sum and Theorem 1 states that the fluctuations of the

random trigonometric sum  $\sum_{k=1}^N \sin 2\pi n_k x$  around its nonrandom average part always have a mixed normal limit distribution. If  $\cup_{k=1}^{\infty} |I_k| = \mathbb{N}$ , i.e. there are no gaps between the blocks  $I_k$ , then  $\sum_{k=1}^n \lambda_k(x) = O(1)$  for any fixed  $x$  and thus (1.3) holds without the  $\lambda_k(x)$ , yielding the result of Bobkov and Götze [2]. Letting  $\Delta_k$  denote the number of integers between  $I_k$  and  $I_{k+1}$  (the "gaps"), we will see that the CLT (1.3) also holds with  $\lambda_k(x) = 0$  if  $\Delta_k$  is nondecreasing and  $\Delta_k = O(k^\gamma)$  for some  $\gamma < 1/4$ . If  $\Delta_k$  grows at least exponentially, then so does the sequence  $(A_k)$ , where  $A_k$  denotes the smallest integer of  $I_k$ . Now

$$\lambda_k(x) = \frac{\sin d\pi x}{d \sin \pi x} \sin 2\pi(A_k + d/2 - 1/2)x \quad (1.5)$$

and from the CLT of Salem and Zygmund [7] it follows that the limit distribution of  $N^{-1/2} \sum_{k=1}^N \lambda_k(x)$  is  $N(0, g^*)$ , where

$$g^*(x) = \frac{\sin^2 d\pi x}{2d^2 \sin^2 \pi x}. \quad (1.6)$$

By Theorem 1, the limit distribution of  $N^{-1/2} \sum_{k=1}^N (\sin 2\pi n_k x - \lambda_k(x))$  is  $N(0, g)$  with  $g$  in (1.4) and the convolution of these two mixed Gaussian laws is  $N(0, 1/2)$ , which is exactly the limit distribution of  $N^{-1/2} \sum_{k=1}^N \sin 2\pi n_k x$  by the theorem of Salem and Zygmund, since  $(n_k)$  grows exponentially. Thus the pure Gaussian limit distribution of  $N^{-1/2} \sum_{k=1}^N \sin 2\pi n_k x$  is obtained as the combination of two mixed Gaussian distributions  $N(0, g)$  with  $g$  in (1.4) and  $N(0, g^*)$  with  $g^*$  in (1.6).

It is worth noting that for any fixed  $x \in (0, 1)$ ,  $\sin 2\pi n_k x - \lambda_k(x)$  are independent, uniformly bounded mean zero random variables on  $(\Omega, \mathcal{A}, \mathbb{P})$  and

$$\begin{aligned} \mathbb{E}(\sin 2\pi n_k x - \lambda_k(x))^2 &= \mathbb{E}(\sin^2 2\pi n_k x) - \lambda_k^2(x) \\ &= \frac{1}{d} \sum_{j \in I_k} \sin^2 2\pi j x - \left( \frac{1}{d} \sum_{j \in I_k} \sin 2\pi j x \right)^2 = g(x) \end{aligned}$$

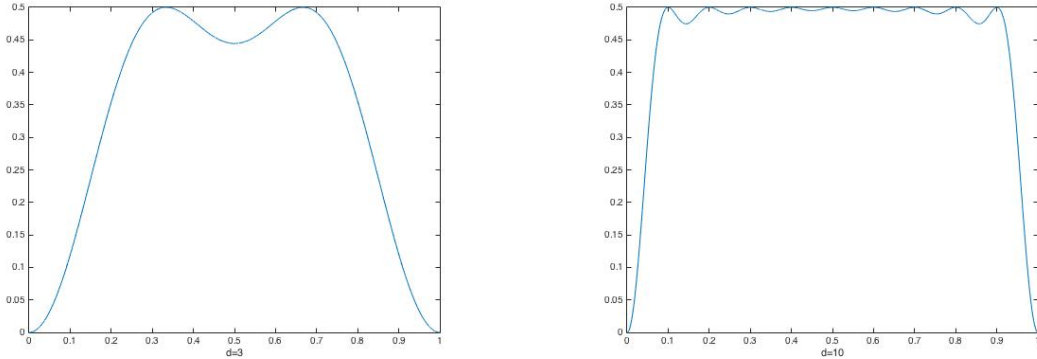
by elementary calculations. Thus by the law of the iterated logarithm we have for any fixed  $x \in (0, 1)$  with  $\mathbb{P}$ -probability 1

$$\limsup_{N \rightarrow \infty} \frac{1}{\sqrt{2N \log \log N}} \sum_{k=1}^N (\sin 2\pi n_k x - \lambda_k(x)) = \sqrt{g(x)}. \quad (1.7)$$

By Fubini's theorem, with  $\mathbb{P}$ -probability 1 relation (1.7) holds for almost every  $x \in (0, 1)$  with respect to Lebesgue measure, yielding the LIL corresponding to (1.3). Actually, the previous argument also shows that for any fixed  $x \in (0, 1)$  we have (1.3) over the probability space  $(\Omega, \mathcal{A}, \mathbb{P})$ , with  $N(0, g)$  replaced by  $N(0, g(x))$ . However, Fubini's theorem does not work for distributional results and thus we cannot interchange the role of  $x \in (0, 1)$  and  $\omega \in \Omega$  and we will need an elaborate argument in Section 2 to prove Theorem 1.

Formula (1.4) shows that for any  $0 < x < 1$  we have  $\lim_{x \rightarrow \infty} g(x) = 1/2$  and thus for large  $d$  the sequence  $\sin 2\pi n_k x - \lambda_k(x)$  nearly satisfies the ordinary CLT and LIL

with limit distribution  $N(0, 1/2)$  and  $\text{lmsup} = 1/2$ , just as lacunary trigonometric series with exponential gaps. Formally, this is not surprising since for large  $d$  the expected gaps  $\mathbb{E}(n_{k+1} - n_k)$  in our sequence are large. As the pictures of  $g$  for  $d = 3$  and  $d = 10$  below show, however, the near CLT and LIL actually hold for relatively small values of  $d$  such as  $d = 10$ . Thus the reason of the near CLT and LIL is not solely large gaps in the the sequence  $(n_k)$  but the random fluctuations of the sequence  $(n_k)$  as well.



The analogue of Theorem 1 is valid also in the case when  $n_1, n_2, \dots$  have continuous uniform distribution over the intervals  $I_1, I_2, \dots$ . To formulate the result, define the probability measure  $\mu$  on the Borel sets of  $\mathbb{R}$  by

$$\mu(A) = \frac{1}{\pi} \int_A \left( \frac{\sin x}{x} \right)^2 dx, \quad A \subset \mathbb{R}.$$

**Theorem 2.** *Let  $n_1, n_2, \dots$  be a sequence of independent random variables on a probability space  $(\Omega, \mathcal{A}, \mathbb{P})$  such that  $n_k$  has continuous uniform distribution on the interval  $[A_k, A_k + B]$ , where  $A_{k+1} - A_k \geq B + 2$ ,  $k = 1, 2, \dots$ . Let  $\lambda_k(x) = \mathbb{E}(\sin n_k x)$ . Then  $\mathbb{P}$ -almost surely*

$$\frac{1}{\sqrt{N}} \sum_{k=1}^N (\sin n_k x - \lambda_k(x)) \xrightarrow{d} F \tag{1.8}$$

with respect to the probability space  $(\mathbb{R}, \mathcal{B}, \mu)$ , where the characteristic function of  $F$  is

$$\phi(\lambda) = \int_{-\infty}^{+\infty} \exp \left( -\frac{\lambda^2}{4} \left( 1 - \frac{4 \sin^2(Bx/2)}{B^2 x^2} \right) \right) d\mu(x). \tag{1.9}$$

## 2 Proofs

We will give the proof of Theorem 2, where the calculations are slightly simpler. Let

$$\varphi_k(x) = \sin n_k x - \mathbb{E}(\sin n_k x)$$

and

$$T_N = \frac{1}{\sqrt{N}} \sum_{k=1}^N \varphi_k(x).$$

By  $A_{k+1} - A_k \geq B + 2$  and the fact that

$$\int_{-\infty}^{+\infty} \cos \alpha x \left( \frac{\sin x}{x} \right)^2 dx = 0 \quad \text{for } |\alpha| > 2 \quad (2.10)$$

(see e.g. Hartman [5]) it follows that for every fixed  $\omega \in \Omega$  the functions  $\varphi_k$  are orthogonal over  $L^2_\mu(\mathbb{R})$  and thus elementary algebra shows that the  $L^2_\mu(\mathbb{R})$  norm of  $|T_M - T_{N^3}|$  is at most  $C/\sqrt{N}$  for  $N^3 \leq M \leq (N+1)^3$  with an absolute constant  $C$ . Hence to prove (1.8) it suffices to show that  $T_{N^3} \xrightarrow{d} F$   $\mathbb{P}$ -a.s.

A simple calculation shows that

$$\begin{aligned} \lambda_k(x) &= \mathbb{E}(\sin n_k x) = \frac{1}{B} \int_{A_k}^{A_k+B} \sin tx dt = \frac{1}{Bx} (\cos A_k x - \cos(A_k + B)x) \\ &= \frac{2 \sin(Bx/2)}{Bx} \sin(A_k + B/2)x \end{aligned} \quad (2.11)$$

and

$$\mathbb{E}(\cos 2n_k x) = \frac{1}{B} \int_{A_k}^{A_k+B} \cos 2tx dt = \frac{\sin Bx}{Bx} \cos(2A_k + B)x.$$

Thus

$$\begin{aligned} \mathbb{E}\varphi_k^2(x) &= \mathbb{E}(\sin^2 n_k x) - \lambda_k^2(x) = \frac{1}{2} (1 - \mathbb{E}(\cos 2n_k x)) - \lambda_k^2(x) \\ &= \frac{1}{2} - \frac{\sin Bx}{2Bx} \cos(2A_k + B)x - \frac{4 \sin^2(Bx/2)}{B^2 x^2} \sin^2(A_k + B/2)x \\ &= \left( \frac{1}{2} - \frac{2 \sin^2(Bx/2)}{B^2 x^2} \right) + \left( \frac{2 \sin^2(Bx/2)}{B^2 x^2} - \frac{\sin Bx}{2Bx} \right) \cos(2A_k + B)x. \end{aligned}$$

From (2.10),  $A_{k+1} - A_k \geq B + 2$  and elementary trigonometric identities it follows that the functions  $\cos(2A_k + B)x$  are orthogonal in  $L^2_\mu(\mathbb{R})$  and thus the Rademacher-Menshov convergence theorem implies that  $\sum_{k=1}^\infty k^{-1} \cos(2A_k + B)x$  converges  $\mu$ -almost everywhere. Consequently, the Kronecker lemma implies

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=1}^N \cos(2A_k + B)x = 0 \quad \mu - \text{a.e.}$$

and thus

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=1}^N \mathbb{E}\varphi_k^2(x) = \frac{1}{2} \left( 1 - \frac{4 \sin^2(Bx/2)}{B^2 x^2} \right) \quad \mu - \text{a.e.}$$

Since for fixed  $x$   $\varphi_k^2(x) - \mathbb{E}\varphi_k^2(x)$ ,  $k = 1, 2, \dots$  are independent, uniformly bounded, zero mean random variables, the strong law of large numbers yields

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=1}^N (\varphi_k^2(x) - \mathbb{E}\varphi_k^2(x)) = 0 \quad \mathbb{P} - \text{a.s.}$$

and thus we conclude that for  $\mu$ -a.e.  $x$  we have  $\mathbb{P}$ -almost surely

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=1}^N \varphi_k^2(x) = \frac{1}{2} \left( 1 - \frac{4 \sin^2(Bx/2)}{B^2 x^2} \right). \quad (2.12)$$

By Fubini's theorem,  $\mathbb{P}$ -almost surely the last relation holds for  $\mu$ -almost all  $x \in \mathbb{R}$ . Fix  $\lambda \in \mathbb{R}$ . Using  $|\varphi_k(x)| \leq 2$  and

$$\exp(z) = (1 + z) \exp\left(\frac{z^2}{2} + o(z^2)\right) \quad z \rightarrow 0$$

we get

$$\exp\left(\frac{i\lambda}{\sqrt{N}}\varphi_k(x)\right) = \left(1 + \frac{i\lambda}{\sqrt{N}}\varphi_k(x)\right) \exp\left(-\frac{\lambda^2\varphi_k^2(x)}{2N} + o\left(\frac{\lambda^2\varphi_k^2(x)}{N}\right)\right)$$

as  $N \rightarrow \infty$ , uniformly in  $x$  and the implicit variable  $\omega \in \Omega$ . Thus the characteristic function

$$\phi_{T_N}(\lambda) = \int_{-\infty}^{\infty} \exp\left(\frac{i\lambda}{\sqrt{N}} \sum_{k=1}^N \varphi_k(x)\right) d\mu(x) = \int_{-\infty}^{\infty} \exp\left(\frac{i\lambda}{\sqrt{N}} \sum_{k=1}^N \varphi_k(x, \omega)\right) d\mu(x)$$

of  $T_N$  with respect to the probability space  $(\mathbb{R}, \mathcal{B}, \mu)$  can be written as

$$\begin{aligned} \phi_{T_N}(\lambda) &= \int_{-\infty}^{+\infty} \prod_{k=1}^N \left(1 + \frac{i\lambda}{\sqrt{N}}\varphi_k(x)\right) \\ &\quad \times \exp\left(-\frac{\lambda^2}{2N} \sum_{k=1}^N \varphi_k^2(x)\right) \frac{1}{\pi} \left(\frac{\sin x}{x}\right)^2 dx. \end{aligned}$$

For simplicity let

$$\hat{g}(x) = \frac{1}{2} \left( 1 - \frac{4 \sin^2(Bx/2)}{B^2 x^2} \right).$$

Using  $1 + x \leq e^x$  and  $|\varphi_k(x)| \leq 2$  we get

$$\begin{aligned} \left| \prod_{k=1}^N \left(1 + \frac{i\lambda}{\sqrt{N}}\varphi_k(x)\right) \right| &= \prod_{k=1}^N \left(1 + \frac{\lambda^2}{N}\varphi_k^2(x)\right)^{1/2} \\ &\leq \exp\left(\frac{\lambda^2}{2N} \sum_{k=1}^N \varphi_k^2(x)\right) \leq e^{2\lambda^2} \end{aligned} \quad (2.13)$$

and thus the dominated convergence theorem and (2.12) imply  $\mathbb{P}$ -almost surely

$$\phi_{T_N}(\lambda) = \int_{-\infty}^{+\infty} \prod_{k=1}^N \left(1 + \frac{i\lambda}{\sqrt{N}}\varphi_k(x)\right) \exp(-\lambda^2 \hat{g}(x)/2) \frac{1}{\pi} \left(\frac{\sin x}{x}\right)^2 dx + o(1).$$

Since the characteristic function  $\phi(\lambda)$  of  $F$  in (1.8) is given by (1.9), to prove that  $T_{N^3} \xrightarrow{d} F$   $\mathbb{P}$ -a.s., it remains to show that letting

$$\Gamma_N = \int_{-\infty}^{+\infty} \left[ \prod_{k=1}^N \left( 1 + \frac{i\lambda}{\sqrt{N}} \varphi_k(x) \right) - 1 \right] \exp(-\lambda^2 g(x)/2) \frac{1}{\pi} \left( \frac{\sin x}{x} \right)^2 dx,$$

we have

$$\Gamma_{N^3} \xrightarrow{\mathbb{P}\text{-a.s.}} 0.$$

Clearly

$$\begin{aligned} \mathbb{E}|\Gamma_N|^2 &= \mathbb{E} \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \left[ \prod_{k=1}^N \left( 1 + \frac{i\lambda}{\sqrt{N}} \varphi_k(x) \right) - 1 \right] \left[ \prod_{k=1}^N \left( 1 - \frac{i\lambda}{\sqrt{N}} \varphi_k(y) \right) - 1 \right] \\ &\quad \times \exp(-\lambda^2 g(x)/2) \exp(-\lambda^2 g(y)/2) d\mu(x) d\mu(y). \end{aligned} \quad (2.14)$$

Now using the independence of the  $\varphi_k$  and  $\mathbb{E}\varphi_k(x) = \mathbb{E}\varphi_k(y) = 0$  we get

$$\begin{aligned} &\mathbb{E} \left[ \prod_{k=1}^N \left( 1 + \frac{i\lambda}{\sqrt{N}} \varphi_k(x) \right) - 1 \right] \left[ \prod_{k=1}^N \left( 1 - \frac{i\lambda}{\sqrt{N}} \varphi_k(y) \right) - 1 \right] \\ &= \mathbb{E} \left[ \prod_{k=1}^N \left( 1 + \frac{i\lambda}{\sqrt{N}} \varphi_k(x) \right) \left( 1 - \frac{i\lambda}{\sqrt{N}} \varphi_k(y) \right) \right] - 1 \\ &= \mathbb{E} \left[ \prod_{k=1}^N \left( 1 + \frac{i\lambda}{\sqrt{N}} \varphi_k(x) - \frac{i\lambda}{\sqrt{N}} \varphi_k(y) + \frac{\lambda^2}{N} \varphi_k(x) \varphi_k(y) \right) \right] - 1 \\ &= \prod_{k=1}^N \left( 1 + \frac{\lambda^2}{N} \Psi_k(x, y) \right) - 1, \end{aligned}$$

where  $\Psi_k(x, y) = \mathbb{E}\varphi_k(x)\varphi_k(y)$ . Thus interchanging the expectation with the double integral in (2.14) we get

$$\begin{aligned} \mathbb{E}|\Gamma_N|^2 &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \left[ \prod_{k=1}^N \left( 1 + \frac{\lambda^2}{N} \Psi_k(x, y) \right) - 1 \right] \times \\ &\quad \times \exp(-\lambda^2 g(x)/2 - \lambda^2 g(y)/2) d\mu(x) d\mu(y) \\ &\leq \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \left| \prod_{k=1}^N \left( 1 + \frac{\lambda^2}{N} \Psi_k(x, y) \right) - 1 \right| d\mu(x) d\mu(y). \end{aligned}$$

Using  $|\Psi_k(x, y)| \leq 4$  and  $|\log(1+x) - x| \leq Cx^2$  for all  $|x| \leq 1$  and some constant  $C > 0$ , one deduces for all sufficiently large  $N$ ,

$$\left| \log \prod_{k=1}^N \left( 1 + \frac{\lambda^2}{N} \Psi_k(x, y) \right) - \sum_{k=1}^N \frac{\lambda^2}{N} \Psi_k(x, y) \right| \leq \frac{16C\lambda^4}{N}.$$

Thus letting

$$G_N(x, y) := \sum_{k=1}^N \frac{\lambda^2}{N} \Psi_k(x, y)$$

we get, using  $G_N(x, y) \leq 4\lambda^2$ , that

$$\prod_{k=1}^N \left( 1 + \frac{\lambda^2}{N} \Psi_k(x, y) \right) = \exp \{ G_N(x, y) + O(\lambda^4/N) \} = 1 + O(|G_N(x, y)|) + O(1/N).$$

Thus

$$\mathbb{E}|\Gamma_N|^2 \leq C_1 \left( \frac{1}{N} + \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} |G_N(x, y)| d\mu(x) d\mu(y) \right) \quad (2.15)$$

for some constant  $C_1$ . In view of  $A_{k+1} - A_k \geq B + 2$  and (2.10), for any  $\lambda_1 \in [A_k, A_k + B]$ ,  $\lambda_2 \in [A_l, A_l + B]$ ,  $k \neq l$ ,  $\sin \lambda_1 x$  and  $\sin \lambda_2 x$  are orthogonal in  $L_\mu^2(\mathbb{R})$ , which implies that  $\varphi_k$  and  $\varphi_l$  are also orthogonal in  $L_\mu^2(\mathbb{R})$ . Since  $\Psi_k(x, y) \Psi_l(x, y) = \mathbb{E} \varphi_k(x) \varphi_l(x) \varphi_k(y) \varphi_l(y)$ , it follows that

$$\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \Psi_k(x, y) \Psi_l(x, y) d\mu(x) d\mu(y) = 0 \quad \text{for } k \neq l$$

and thus by the Cauchy-Schwarz inequality the last integral in (2.15) is  $O(N^{-1/2})$ . Hence  $\mathbb{E}|\Gamma_N|^2 = O(N^{-1/2})$  and thus  $\sum_{N \in \mathbb{N}} \mathbb{E}|\Gamma_{N^3}|^2 < \infty$ , implying  $\sum_{N \in \mathbb{N}} |\Gamma_{N^3}|^2 < \infty$  and  $\Gamma_{N^3} \rightarrow 0$   $\mathbb{P}$ -a.s., completing the proof of (1.8).

In conclusion we prove the claim made after Theorem 1, namely that if the size of the gaps  $\Delta_k$  between the blocks  $I_k$  is nondecreasing and satisfies

$$\Delta_k = O(k^\gamma), \quad \gamma < 1/4 \quad (2.16)$$

then

$$N^{-1/2} \sum_{k=1}^N \lambda_k(x) \longrightarrow 0 \quad \text{a.s.}$$

and thus (1.3) holds with  $\lambda_k(x) = 0$ . Since we proved our main limit theorem in the continuous case of Theorem 2, we prove our claim also in the context of Theorem 2 in which case we also assume that the intervals  $[A_k, A_k + B]$  have integer endpoints. In view of (2.11) it suffices to show that

$$N^{-1/2} \sum_{k=1}^N e^{iA_k x} \longrightarrow 0 \quad \text{a.s.} \quad (2.17)$$

and here nothing changes if we replace  $x$  by  $2\pi x$ . In the case of constant  $\Delta_k$  we have  $A_k = Dk + D^*$  for some constants  $D > 0$  and  $D^*$  and (2.17) is obvious by an explicit



computation of the sum. Thus we can assume  $\Delta_k \uparrow \infty$ , and then also  $A_{k+1} - A_k \uparrow \infty$ . Recalling that the  $A_k$  are integers, let us break the sum  $\sum_{k=1}^N e^{2\pi i A_k x}$  into subsums

$$Z_{N,r} = \sum_{k \leq N, A_{k+1} - A_k = r} e^{2\pi i A_k x}, \quad r = 1, 2, \dots \quad (2.18)$$

Clearly  $Z_{N,r}$  consists of  $M_r$  consecutive terms of  $\sum_{k=1}^N e^{2\pi i A_k x}$  for some  $M_r \geq 0$  and thus in the case  $M_r \geq 1$  we have for some integer  $P_r \geq 0$ ,

$$|Z_{N,r}| = \left| \sum_{j=0}^{M_r-1} e^{2\pi i (P_r + jr)x} \right| = \left| \sum_{j=0}^{M_r-1} e^{2\pi i jr x} \right| \leq \frac{1}{|e^{2\pi i r x} - 1|} \leq \frac{C}{\langle rx \rangle},$$

except when  $rx$  is an integer, where  $C$  is an absolute constant and  $\langle t \rangle$  denotes the distance of  $t$  from the nearest integer. From a well known result in Diophantine approximation theory (see e.g. Kuipers and Niederreiter [6], Definition 3.3. on p. 121 and Exercise 3.5 on page 130), for every  $\varepsilon > 0$  and almost all  $x$  in the sense of Lebesgue measure we have  $\langle nx \rangle \geq cn^{-(1+\varepsilon)}$  for some constant  $c = c(x) > 0$  and all  $n \geq 1$ . This shows that  $Z_{N,r} = O(r^{1+\varepsilon})$  a.e. and since by (2.16) the largest  $r$  actually occurring in breaking  $\sum_{k=1}^N e^{2\pi i A_k x}$  into a sum of  $Z_{N,r}$ 's is at most  $C_1 N^\gamma$ , we have

$$\left| \sum_{k=1}^N e^{2\pi i A_k x} \right| \leq C_2 \sum_{r \leq C_1 N^\gamma} r^{1+\varepsilon} = o(\sqrt{N}) \quad \text{a.e.}$$

by  $\gamma < 1/4$ , upon choosing  $\varepsilon$  small enough.

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